

ECE 7251: Signal Detection and Estimation

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Lecture 26, 3/15/01:
Uniformly Most Powerful Tests

An Introductory Case

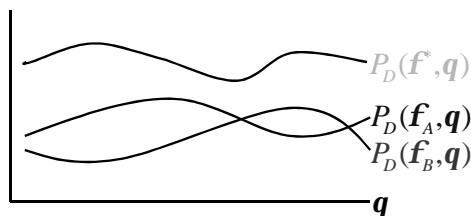
- Usual parametric data model $p(y; \mathbf{q})$
- Consider a composite problem:

$$H_0 : \mathbf{q} = \mathbf{q}_0, H_1 : \mathbf{q} \in \mathcal{S}_1$$
- A test \mathbf{f}^* is uniformly most powerful of level $\mathbf{a} = P_{FA}$ if it has a better P_D (or at least as good as) than any other \mathbf{a} -level test

$$P_D(\mathbf{f}^*; \mathbf{q}) = E_{\mathbf{q}}[\mathbf{f}^*] \geq E_{\mathbf{q}}[\mathbf{f}] = P_D(\mathbf{f}, \mathbf{q})$$
for all $\mathbf{q} \in \mathcal{S}_1$, for all \mathbf{f}
- Hero uses notation \mathbf{b} instead of P_D

Graphical Interpretation

$$\mathbf{a} = P_{FA}(\mathbf{f}_B) = P_{FA}(\mathbf{f}_A) = P_{FA}(\mathbf{f}^*)$$



Finding UMP Tests (When They Exist)

- Find the most powerful \mathbf{a} -level (recall $\mathbf{a} = P_{FA}$) test for a fixed \mathbf{q}
 - Just the Neyman-Pearson test
- If the decision regions do not vary with \mathbf{q} , then the test is UMP

Gaussian Mean Example

- Suppose we have n i.i.d samples

$$y_i \sim \mathcal{N}(\mathbf{m}, \mathbf{s}^2)$$

- Assume \mathbf{s}^2 is known, but \mathbf{m} is not
- Consider three cases

$$H_0 : \mathbf{m} = 0$$

$$\text{Case I: } H_1 : \mathbf{m} > 0$$

$$\text{Case II: } H_1 : \mathbf{m} < 0$$

$$\text{Case III: } H_1 : \mathbf{m} \neq 0$$

Sufficies to use

$$\bar{y} = \frac{1}{n} \sum_{i=1}^n y_i$$

$$\bar{y} \sim \mathcal{N}(\mathbf{m}, \mathbf{s}^2/n)$$

The Gaussian Likelihood Ratio

$$\begin{aligned} \Lambda(y; \mathbf{m}) &= \frac{p(y; \mathbf{m})}{p(y; 0)} = \frac{\exp[-(\bar{y} - \mathbf{m})^2 / (2\mathbf{s}^2/n)]}{\exp[-\bar{y}^2 / (2\mathbf{s}^2/n)]} \\ &= \frac{\exp[-\bar{y}^2 + 2\bar{y}\mathbf{m} - \mathbf{m}^2] / (2\mathbf{s}^2/n)}{\exp[-\bar{y}^2 / (2\mathbf{s}^2/n)]} \\ &= \exp \left[\frac{n\mathbf{m}}{\mathbf{s}^2} \bar{y} - \frac{n\mathbf{m}^2}{2\mathbf{s}^2} \right] \stackrel{H_1}{\underset{H_0}{\geq}} t \\ &= \frac{\sqrt{n}\mathbf{m}}{\mathbf{s}} \bar{y} \stackrel{H_1}{\underset{H_0}{\geq}} \left(\ln t + \frac{n\mathbf{m}^2}{2\mathbf{s}^2} \right) \frac{\mathbf{s}}{\sqrt{n}} \equiv g \end{aligned}$$

Case I: $m > 0$

$$\frac{\sqrt{nm}}{s} \bar{y} \underset{H_0}{\geq} g \longrightarrow \frac{\sqrt{n}}{s} \bar{y} \underset{H_0}{\geq} gm = g^+$$

- Set the threshold to get the right “level”

$$a = P_{FA} = \Pr \left[\frac{\sqrt{n}}{s} \bar{y} > g^+ \mid H_0 \right] = Q(g^+)$$

$$g^+ = Q^{-1}(a)$$

- Notice the test does not depend on m hence, it is UMP

Power of the Single-Sided Test (Case I)

$$P_D = \Pr \left[\frac{\sqrt{n}}{s} \bar{y} > g^+ \mid H_1 \right]$$

$$= \Pr \left[\frac{\sqrt{n}}{s} (\bar{y} - m) > g^+ - m \frac{\sqrt{n}}{s} \mid H_1 \right]$$

$$= Q \left(g^+ - \frac{\sqrt{n}}{s} m \right)$$

$$= Q \left(Q^{-1}(a) - \frac{\sqrt{n}}{s} m \right) \equiv Q(Q^{-1}(a) - d)$$

Case II: $m < 0$

$$\frac{\sqrt{nm}}{s} \bar{y} \underset{H_0}{\geq} g \longrightarrow \frac{\sqrt{n}}{s} \bar{y} \underset{H_1}{\geq} gm = g^-$$

Notice flip!

$$a = P_{FA} = \Pr \left[\frac{\sqrt{n}}{s} \bar{y} < g^- \mid H_0 \right] = 1 - Q(g^-)$$

$$g^- = Q^{-1}(1 - a)$$

- Case II is UMP also

Power of the Single-Sided Test (Case II)

$$P_D = \Pr \left[\frac{\sqrt{n}}{s} \bar{y} < g^- \mid H_1 \right]$$

$$= \Pr \left[\frac{\sqrt{n}}{s} (\bar{y} - m) < g^- - m \frac{\sqrt{n}}{s} \mid H_1 \right]$$

$$= 1 - Q \left(g^- - \frac{\sqrt{n}}{s} m \right) = 1 - Q \left(Q^{-1}(1 - a) - \frac{\sqrt{n}}{s} m \right)$$

$$= 1 - Q(Q^{-1}(1 - a) - d)$$

Case III: $m \neq 0$

$$\frac{\sqrt{nm}}{s} \bar{y} \underset{H_0}{\geq} g$$

- Uh oh... we can't just absorb m into the threshold anymore without effecting the inequalities!
- Decision region varies with sign of m
- No UMP test exists

Cauchy Median Example

- Suppose we have a single sample from the density

$$p(y; \mathbf{q}) = \frac{1}{p} \frac{1}{1 + (y - \mathbf{q})^2}$$

and we want to decide $H_0 : \mathbf{q} = 0, H_1 : \mathbf{q} > 0$

- Likelihood ratio is

$$\frac{p(y; \mathbf{q})}{p(y; 0)} = \frac{1 + y^2}{1 + (y - \mathbf{q})^2} \underset{H_0}{\geq} t$$

- Decision region depends on θ , so no UMP exists!

The Monotone Likelihood Ratio Condition

$$H_0 : \mathbf{q} = \mathbf{q}_0, H_1 : \mathbf{q} \in \mathcal{S}_1$$

- Suppose we have a Fisher Factorization

$$p(y; \mathbf{q}) = g(T(y), \mathbf{q})h(y)$$

- A UMP test of any level α exists if the likelihood ratio is either monotone increasing or decreasing in T for all $\mathbf{q} \in \mathcal{S}_1$

$$\Lambda(y; \mathbf{q}) = \frac{p(y; \mathbf{q})}{p(y; \mathbf{q}_0)} = \frac{g(T, \mathbf{q})}{g(T, \mathbf{q}_0)} \equiv \Lambda(T; \mathbf{q})$$

↑
(abuse)

Densities Satisfying Monotone Likelihood Ratio Condition

- Suppose we have a one-sided test:

$$H_0 : \mathbf{q} = \mathbf{q}_0, H_1 : \mathbf{q} > \mathbf{q}_0$$

$$\text{or } H_0 : \mathbf{q} = \mathbf{q}_0, H_1 : \mathbf{q} < \mathbf{q}_0$$

- The following satisfy the MLR condition:
 - i.i.d. samples from 1-D exponential family (Gaussian, Bernoulli, Exponential, Poisson, Gamma, Beta)
 - i.i.d. samples from uniform density $U(0, \mathbf{q})$
 - i.i.d. samples from shifted Laplace
- also works if $H_0 : \mathbf{q} < \mathbf{q}_0$

Densities *Not* Satisfying Monotone Likelihood Ratio Condition

- Gaussian with single-sided H_1 on mean but *unknown* variance
- Cauchy density with single-sided H_1 on centrality parameter
- Exponential family with double-sided H_1