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# Discussion of Market Expectations in the Cross Section of Present Values

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Market Expectations

## The problem tackled by this paper

- Market return

$$r_{mt+1} = \mu_t + \eta_{t+1}$$

- Present-value decomposition of market M/B ratio

$$x_{mt} \approx a - b_\mu \mu_t + b_g g_t$$

with VAR(1) dynamics of book ROE,  $g_t$ , and expected return,  $\mu_t$

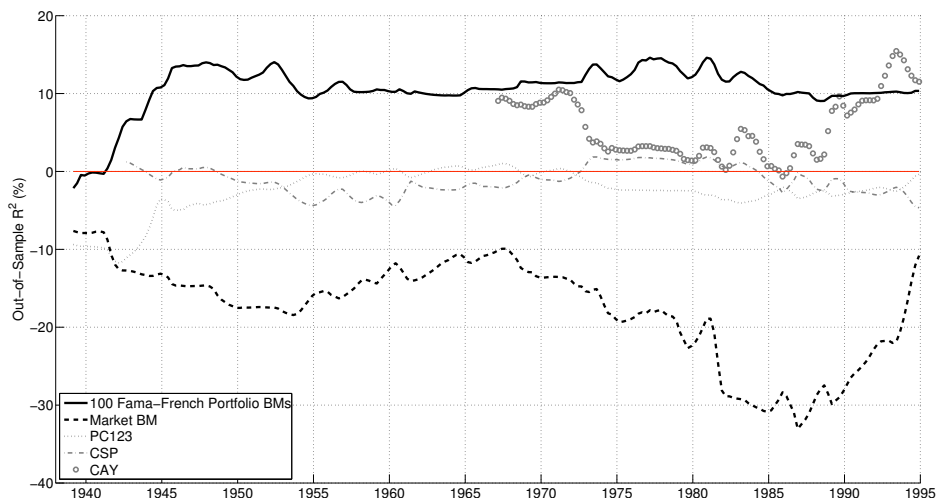
- Presence of  $g_t$  obscures “signal”  $\mu_t$
- This paper: Cross-sectional information helps filter out  $\mu_t$  to improve predictive regressions
- Much of my discussion focuses on Kelly and Pruitt (2012, “3PRF”, WP), the paper that supplies the methodology

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# The results: U.S.

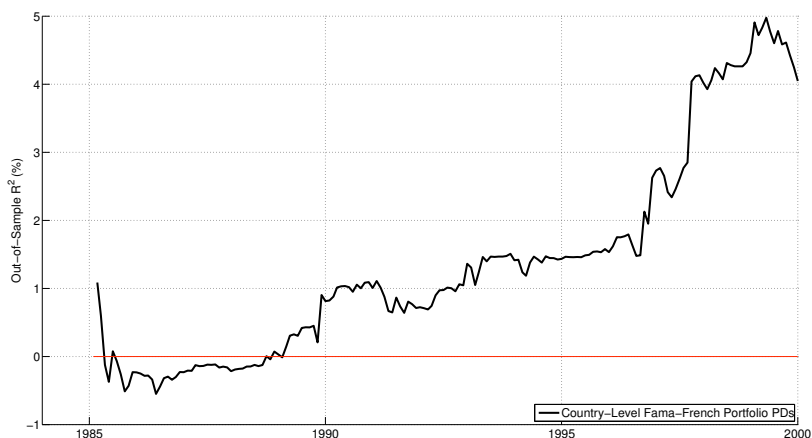
Figure 1: OUT-OF-SAMPLE  $R^2$  BY SAMPLE SPLIT DATE, ONE YEAR RETURNS



Notes: Out-of-sample  $R^2$  across sample split dates. Forecasts are based on a single PLS factor from 100 book-to-market ratios of size and value-sorted portfolios, the aggregate book-to-market ratio, the cross-section premium of Polk et al. (2006), the consumption-wealth ratio of Lettau and Ludvigson (2001), and the first three principal components of the 100 book-to-market ratio cross section.

# The results: International

Figure 3: OUT-OF-SAMPLE  $R^2$  BY SAMPLE SPLIT DATE, ONE MONTH INTERNATIONAL RETURNS



Notes: Out-of-sample percentage  $R^2$  across sample split dates for forecasts of one month international stock returns using a single PLS factor from 42 price-dividend ratios of high value and low value portfolios across 21 countries (Fama and French (1998)). See Section III.A.5 for list of countries.

## How does it work? Example

- Define:  $g_t \equiv$  ROE component orthogonal to expected returns
- Assume: Three-asset cross-section where M/B ratios load on  $\mu_t$  and  $g_t$  with cross-sectionally **uncorrelated** loadings

$$\begin{aligned}x_{1t} &= 2\mu_t - 1.5g_t & \tilde{x}_{1t} &= \mu_t - 0.5g_t \\x_{2t} &= \mu_t & \tilde{x}_{2t} &= g_t \\x_{3t} &= -1.5g_t & \tilde{x}_{3t} &= -\mu_t - 0.5g_t\end{aligned}$$

- First-stage t.s. regressions of  $x_{it}$  on  $r_{mt+1}$  (large  $T$ ): slopes proportional to

$$\begin{aligned}\phi_1 &= 2 & \tilde{\phi}_1 &= 1 \\ \phi_2 &= 1 & \tilde{\phi}_2 &= 0 \\ \phi_3 &= 0 & \tilde{\phi}_3 &= -1\end{aligned}$$

- Second-stage c.s. regressions of  $\tilde{x}_{it}$  on  $\tilde{\phi}_i$  each  $t$ : slopes

$$F_t = \text{const.} \times [(\mu_t - 0.5g_t) + 0 - (-\mu_t - 0.5g_t)] = \text{const.} \times \mu_t$$

## How does it work? Crucial assumption

- Assumption that loadings on  $\mu_t$  and  $g_t$  are c.s. **uncorrelated** is important: Second stage regression slopes  $F_t$  are then...
  - “Long” in assets with positive M/B loadings on  $\mu_t$
  - “Short” in assets with negative M/B loadings on  $\mu_t$
  - “Long” and “short” in assets with M/B that load similarly on  $g_t$ :  $g_t$  exposure cancels out
  - As  $N \rightarrow \infty$ , sample c.s. correlation closer to zero population c.s. correlation:  $\mu_t$  consistently estimated
- But what if loadings on  $\mu_t$  and  $g_t$  are c.s. **correlated**?
  - Lucky: For stock market application in JF paper, correlation seems to be close to zero
  - But this may not be true in other applications
  - Remedy: Use proxies for  $g_t$  in addition to  $\mu_t$  proxy
  - But that means we have to take a stand on **all** of the systematic factors driving  $M/B$  ratios

## Concern: Large $N$ , small $T$

- What are the properties of the estimator under the null hypothesis of no predictability?
- Simulation: No-predictability & pure-noise M/B null

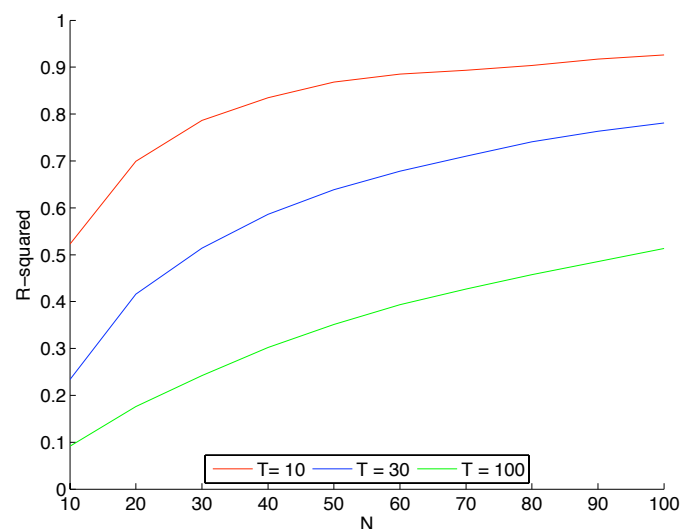
$$r_{mt+1} = \eta_{t+1}$$

$$x_t = \epsilon_t$$

where

$$\begin{pmatrix} \eta_{t+1} \\ \epsilon_t \end{pmatrix} \sim \mathcal{N}(0, I_{N+1})$$

## Large $N$ , small $T$ : Spurious fit as $N$ grows

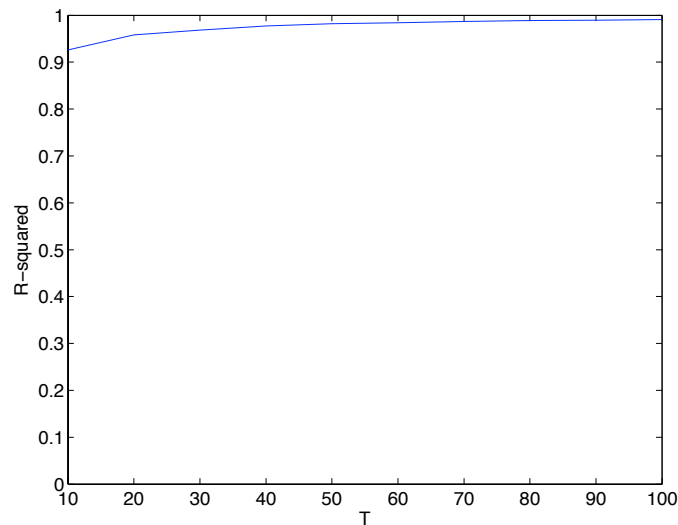


Mean third-stage R-squared under no-predictability null  
(100 simulations for each  $(N, T)$  pair)

# Large $N$ , small $T$ : Example with $T = 3$ and $N = 9$

Pure noise returns	$r_{mt+1}$	-1.35	0.91	-1.49	
	$x_{it}$	-0.52	-1.26	0.39	$\hat{\phi}_i$
		-0.82	-0.44	-0.22	-5.56
Pure noise M/B		-0.20	1.07	2.51	-1.73
		-0.93	-2.62	0.34	-0.49
		0.40	-0.06	-1.14	-0.20
		1.94	0.43	2.08	-0.02
		-1.89	-0.96	-1.83	1.18
		0.44	-0.82	1.51	1.83
		-0.68	-0.37	-0.07	10.88
	$\hat{F}_t$	0.0044	0.0100	0.0040	18.28
	$\hat{y}_{t+1}$	-0.98	0.80	-1.75	
	$R^2$				
		0.94			

## $N$ and $T$ grow simultaneously: Spurious fit with $N = T^2$



Mean third-stage R-squared under no-predictability null with  $N = T^2$   
 (100 simulations for each  $(N, T)$  pair)

## Spurious predictability

- Thus, when  $N$  is not small relative to  $T$ , there is a bias that overstates predictability
- It seems that this is not just a small-sample bias: also asymptotic bias if  $N$  grows sufficiently fast relative to  $T$ 
  - Theorem 1 in Kelly and Pruitt (2012, “3PRF”) seems to need additional assumption:  $N$  cannot grow too fast relative to  $T$
- Not a concern for the empirical results in the JF paper, as out-of-sample tests are not affected by this bias
- But concern underscores importance of out-of-sample testing
- Calls for further study of small-sample and asymptotic properties of this estimator

## Summing up

- Nice idea
- Impressive empirical results: Strong out-of-sample predictability of stock market returns
- Some further work necessary on the properties of the 3PRF estimator
  - Correlated factor loadings
  - Small- $T$ /large- $N$  behavior
  - Asymptotic behavior when  $N$  grows fast relative to  $T$
- These concerns do not affect out-of-sample tests: results in the JF paper are robust